

Bayesian Estimation of DSGE Models: Lessons from Second-order Approximations

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Motivation

- Dynamic Stochastic General Equilibrium (DSGE) model
- Likelihood-based estimation with log-linearization:
DYNARE!!
- For welfare analysis, first-order accurate approximation is not sufficient: (Kim and Kim, 2003)

Example: Welfare Reversal

- Two countries, $i = 1, 2$
- No dynamics
- Preference: CRRA

$$u(C_i) = \frac{C_i^{1-\tau} - 1}{1 - \tau}$$

- Technology: stochastic endowment

$$Z_i = \exp(u_i)$$

where u_i is iid with

$$u_i = \begin{cases} 1 & \text{with prob. } 1/2 \\ -1 & \text{with prob. } 1/2 \end{cases}$$

Example: Welfare Reversal

- Autarchy: $C_i^A = Z_i$
- Complete market, risk-sharing: $C_i^M = \frac{Z_1 + Z_2}{2}$
- Deterministic steady-state: $\bar{Z} = \bar{C} = 1$
- Log-linearization:

$$\begin{aligned}\hat{c}_i^{M1} &= \frac{1}{2}(\hat{z}_1 + \hat{z}_2) \\ C_i^{M1} &= \sqrt{Z_1 Z_2}\end{aligned}$$

- Second-order approximation:

$$\begin{aligned}\hat{c}_i^{M2} &= \frac{1}{2}(\hat{z}_1 + \hat{z}_2) + \frac{1}{8}(\hat{z}_1 - \hat{z}_2)^2 \\ C_i^{M2} &= C_i^{M1} \exp\left(\frac{1}{8}(\log Z_1 - \log Z_2)^2\right)\end{aligned}$$

Example: Welfare Reversal

- Welfare: Expected utility

k	$\mathbf{E} [u (C_i^k)]$	$\tau = 0.5$	$\tau = 2$
$A :$	$\frac{1}{2}u(e) + \frac{1}{2}u(e^{-1})$.2553	-.5431
$M :$	$\frac{1}{4}u(e) + \frac{1}{4}u(e^{-1}) + \frac{1}{2}u\left(\frac{e+e^{-1}}{2}\right)$.3698	-.0956
$M1 :$	$\frac{1}{4}u(e) + \frac{1}{4}u(e^{-1}) + \frac{1}{2}u(1)$.1276	-.2715
$M2 :$	$\frac{1}{4}u(e) + \frac{1}{4}u(e^{-1}) + \frac{1}{2}u\left(e^{1/2}\right)$.4117	-.0748

Motivation

- Second-order approximation: DYNARE!! Again!!
- Likelihood-based estimation of second-order approximated models is not trivial.
- Short cut:
 - (1) Estimation based on linearized model
 - (2) Plugging the linear estimates into the second-order approximated model to do policy analysis
- Can this practice be misleading?

New Keynesian DSGE Model

- Representative household
- Monopolistically competitive firms

$$Y_t = \left(\int_0^1 Y_t(j)^{1-\nu} dj \right)^{\frac{1}{1-\nu}}$$

- Nominal rigidities: quadratic adjustment cost

$$AC_t(j) = \frac{\varphi}{2} \left(\frac{P_t(j)}{P_{t-1}(j)} - \pi \right)^2 Y_t(j)$$

- No capital accumulation
- Linear production with perfectly competitive factor market

New Keynesian DSGE Model

- Monetary policy:

Interest feedback rule with nominal interest rate target

$$R_t = R_t^{*1-\rho_R} R_{t-1}^{\rho_R} \exp(\epsilon_{R,t})$$

$$R_t^* = r\pi^* \left(\frac{\pi_t}{\pi^*}\right)^{\psi_1} \left(\frac{Y_t}{Y_t^*}\right)^{\psi_2}$$

- Fiscal policy:

Passive with exogenously driven government spending

- Three shocks:

Technology, government spending, and monetary policy

Symmetric Equilibrium

After detrending, in log-deviation (from steady-state) form:

$$1 = \mathbf{E}_t \left[e^{-\tau(\hat{c}_{t+1} - \hat{c}_t) + \hat{R}_t - \hat{\pi}_{t+1} - \hat{z}_{t+1}} \right]$$

$$\begin{aligned} 0 &= \frac{1-\nu}{\nu\varphi\pi^2} \left(1 - e^{\tau\hat{c}_t} \right) + \left(e^{\hat{\pi}_t} - 1 \right) \left[\left(1 - \frac{1}{2\nu} \right) e^{\hat{\pi}_t} + \frac{1}{2\nu} \right] \\ &\quad - \beta \mathbf{E}_t \left[\left(e^{\hat{\pi}_{t+1}} - 1 \right) e^{-\tau(\hat{c}_{t+1} - \hat{c}_t) + (\hat{y}_{t+1} - \hat{y}_t) + \hat{\pi}_{t+1}} \right] \end{aligned}$$

$$e^{\hat{c}_t - \hat{y}_t} = e^{-\hat{g}_t} - \frac{\varphi\pi^2 g}{2} \left(e^{\hat{\pi}_t} - 1 \right)^2$$

$$\hat{R}_t = \rho_R \hat{R}_{t-1} + (1 - \rho_R) \psi_1 \hat{\pi}_t + (1 - \rho_R) \psi_2 (\hat{y}_t - \hat{g}_t) + \epsilon_{R,t}$$

$$\hat{g}_t = \rho_g \hat{g}_{t-1} + \epsilon_{g,t}$$

$$\hat{z}_t = \rho_z \hat{z}_{t-1} + \epsilon_{z,t}$$

Representation

The solution of the rational expectations system takes the form

$$s_t = \Phi(s_{t-1}, \epsilon_t; \theta)$$

where

$$s_t = [\hat{y}_t, \hat{c}_t, \hat{\pi}_t, \hat{R}_t, \epsilon_{R,t}, \hat{g}_t, \hat{z}_t]'$$

$$\epsilon_t = [\epsilon_{R,t}, \epsilon_{g,t}, \epsilon_{z,t}]'$$

Log-linearization and Rational Expectation

- Log-linearized economy

$$\begin{aligned}\hat{c}_t &= \mathbf{E}_t [\hat{c}_{t+1}] - \frac{1}{\tau} (\hat{R}_t - \mathbf{E}_t [\hat{\pi}_{t+1}] - \mathbf{E}_t [\hat{z}_{t+1}]) \\ \hat{\pi}_t &= \beta \mathbf{E}_t [\hat{\pi}_{t+1}] + \kappa \hat{c}_t \\ \hat{y}_t &= \hat{c}_t + \hat{g}_t\end{aligned}$$

where $\kappa = \tau \frac{1-\gamma}{\gamma \pi^2 \varphi}$.

Without rigidities where $\gamma = 0$ and $\varphi = 0$, κ tends to infinity.

- The solution of the log-linearized rational expectations system

$$s_t = \Phi^{(s)} s_{t-1} + \Phi^{(\epsilon)} \epsilon_t$$

Second-Order Approximation

- Judd (1998), Collard and Juillard (2001),
Jin and Judd (2002), Schmitt-Grohe and Uribe (2004),
Kim, Kim, Schaumburg, and Sims (2005),
Swanson, Anderson, and Levin (2005), Klein (2005)
- Partition s_t into endogenous and exogenous variables,
 $[x'_t, z'_t]'$
- The solution of the second-order approximated rational expectations system

$$\begin{aligned}z_t &= \Psi^{(z)} z_{t-1} + \epsilon_t \\x_{j,t} &= \Psi_j^{(1)} w_t + \left(\Psi_j^{(0)} + w_t' \Psi_j^{(2)} w_t \right)\end{aligned}$$

where $w_t = [x'_{t-1}, z'_t]'$.

State-Space Representation: Transition

- Linear transition from log-linearization:

$$s_t = \Phi^{(s)} s_{t-1} + \Phi^{(\epsilon)} \epsilon_t$$

where $s_t = [x'_t, z'_t]'$

- Quadratic transition from second-order approximation:

$$\begin{aligned} z_t &= \Psi^{(z)} z_{t-1} + \epsilon_t \\ x_{j,t} &= \Psi_j^{(1)} w_t + \left(\Psi_j^{(0)} + w_t' \Psi_j^{(2)} w_t \right) \end{aligned}$$

where $w_t = [x'_{t-1}, z'_t]'$

State-Space Representation: Measurement

- Output growth rate, inflation, and nominal interest rate

$$\text{YGR}_t = \gamma^{(Q)} + 100(\hat{y}_t - \hat{y}_{t-1} + \hat{z}_t) + u_{y,t}$$

$$\text{INFL}_t = \pi^{(A)} + 400\hat{\pi}_t + u_{\pi,t}$$

$$\text{INT}_t = \pi^{(A)} + r^{(A)} + 4\gamma^{(Q)} + 400\hat{R}_t + u_{r,t}$$

- Linear measurement:

$$y_t = \Theta^{(0)} + \Theta^{(s)} s_t + u_t$$

Filter and Likelihood

- *Initialization at the beginning of t :* $p(s_{t-1}|Y^{t-1})$
- *Prediction/Forecasting:*

$$p(s_t|Y^{t-1}) = \int p(s_t|s_{t-1}) p(s_{t-1}|Y^{t-1}) ds_{t-1}$$

- *Filtering/Updating by Bayes' Theorem:*

$$p(s_t|Y^t) = \frac{p(y_t|s_t) p(s_t|Y^{t-1})}{p(y_t|Y^{t-1})}$$

- *Likelihood evaluation:*

$$\ln \mathcal{L}(\theta|Y) = \ln p(y_1|\theta) + \sum_{t=2}^T \ln p(y_t|Y^{t-1}, \theta)$$

Particle Filter

- Use Monte Carlo method to integrate out the latent variables
- Pioneering works:
 - Gordon, Salmon, and Smith (1993)
 - Kitagawa (1996)
- Economics literature:
 - Stochastic volatility:
Pitt and Shephard (1999), Kim, Shephard, and Chib (1998),
Duan and Fulop (2007)
 - Neoclassical growth model:
Fernandez-Villaverde and Rubio-Ramirez (2004,2005,2006)

Prior Specification

Parameter	Density	Mean	Std.	Description
τ	Gamma	2	0.5	inverse EIS
ν	Beta	0.1	0.05	degree of imperfect competition
κ	Gamma	0.3	0.1	Phillips curve, $\kappa = \frac{\tau}{\pi^2 \varphi} \left(\frac{1 - \nu}{\nu} \right)$
c/y	Beta	0.85	0.05	consumption-output ratio
ψ_1	Gamma	1.5	0.25	responsiveness to inflation
ψ_2	Gamma	0.5	0.25	responsiveness to output gap
ρ_R	Beta	0.5	0.2	policy inertia
$r^{(A)}$	Gamma	0.55	0.25	steady state real interest rate
$\pi^{(A)}$	Gamma	3	2	target inflation
$\gamma^{(Q)}$	Normal	0.8	0.25	steady state output growth rate

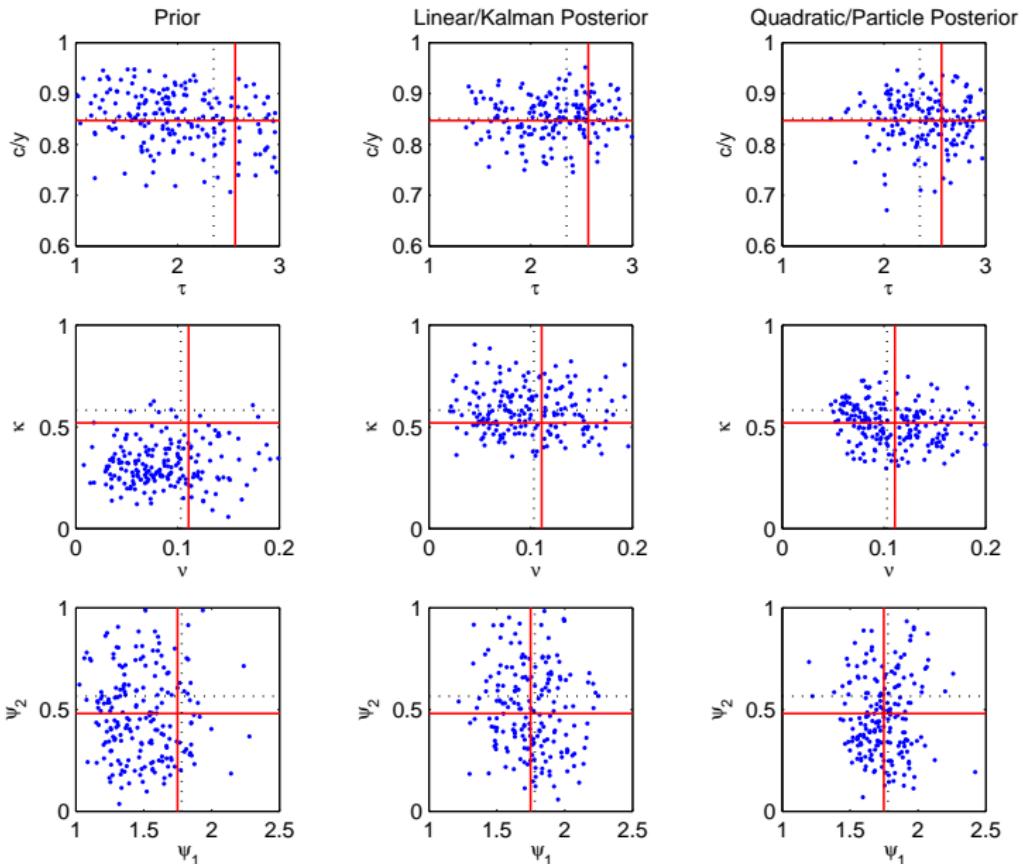
$\rho_g, \rho_z, \sigma_R, \sigma_g$, and σ_z

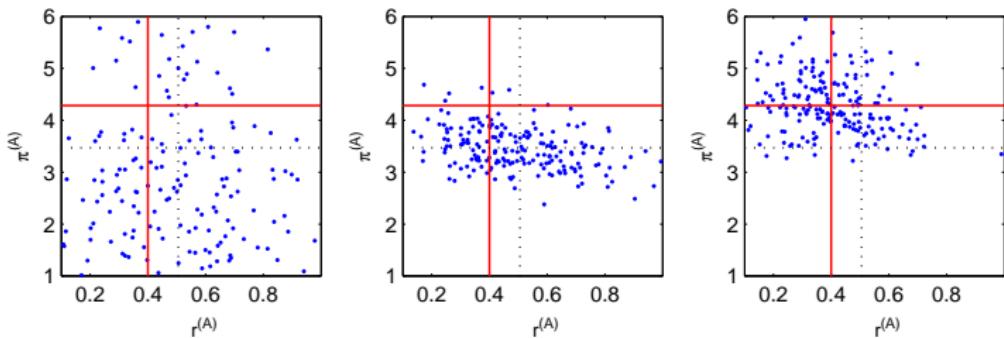
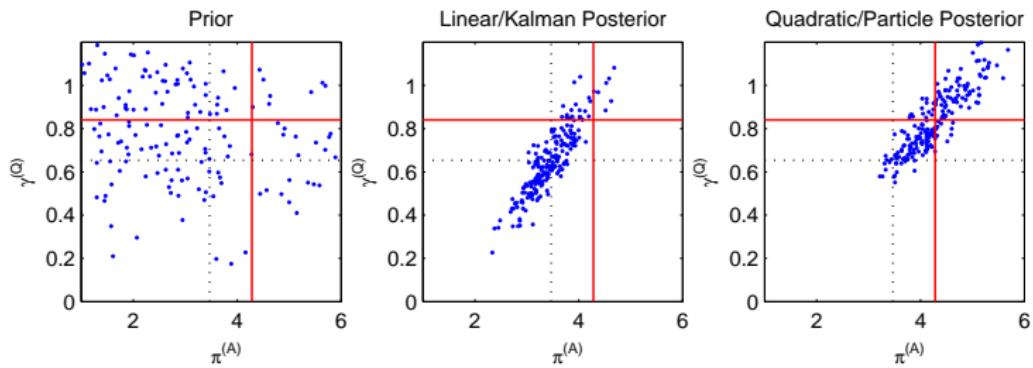
Posterior Distribution

- Log marginal data densities:
Modified harmonic mean estimator by Geweke (1999)

Linear/Kalman	Quadratic/Particle
-323.984	-322.200

- Information on v and c/y
- Steady state parameters, $r^{(A)}$, $\pi^{(A)}$, and $\gamma^{(Q)}$
- Policy parameters, ψ_1 and ψ_2



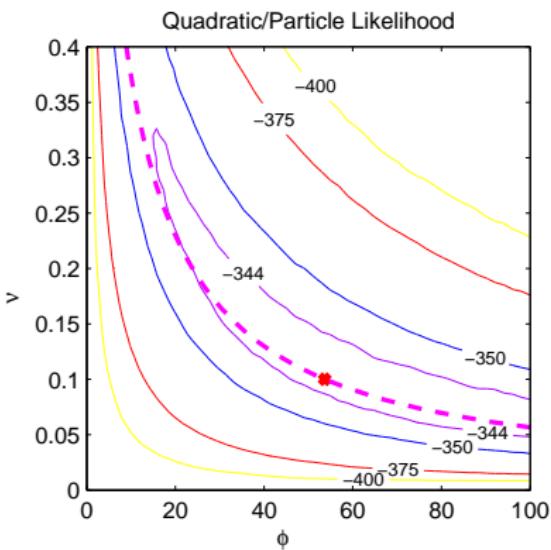
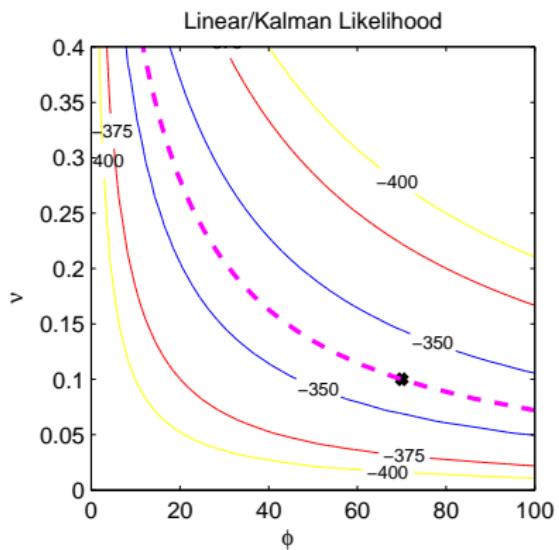


Posterior Distribution

	Prior		Posterior			
	Mean	90% Interval	Linear/Kalman		Quadratic/Particle	
			Mean	90% Interval	Mean	90% Interval
τ	2.004	[1.175, 2.793]	2.352	[1.486, 3.128]	2.565	[1.757, 3.269]
ν	0.100	[0.022, 0.175]	0.103	[0.025, 0.179]	0.111	[0.052, 0.171]
κ	0.300	[0.138, 0.452]	0.582	[0.389, 0.763]	0.520	[0.361, 0.678]
c/y	0.850	[0.771, 0.931]	0.851	[0.773, 0.930]	0.847	[0.776, 0.994]
ψ_1	1.508	[1.091, 1.907]	1.779	[1.417, 2.147]	1.749	[1.456, 2.027]
ψ_2	0.500	[0.110, 0.866]	0.565	[0.123, 0.978]	0.480	[0.140, 0.816]
ρ_R	0.499	[0.148, 0.823]	0.808	[0.760, 0.863]	0.806	[0.755, 0.853]
$r^{(A)}$	0.548	[0.161, 0.918]	0.506	[0.186, 0.797]	0.400	[0.159, 0.632]
$\pi^{(A)}$	2.992	[0.225, 5.765]	3.470	[2.772, 4.205]	4.282	[3.430, 5.212]
$\gamma^{(Q)}$	0.801	[0.380, 1.197]	0.653	[0.405, 0.913]	0.840	[0.630, 1.081]
ρ_g	0.801	[0.651, 0.959]	0.956	[0.927, 0.987]	0.959	[0.932, 0.985]
ρ_z	0.661	[0.429, 0.909]	0.937	[0.905, 0.969]	0.943	[0.917, 0.976]
$100\sigma_R$	0.500	[0.210, 0.790]	0.195	[0.165, 0.229]	0.205	[0.169, 0.242]
$100\sigma_g$	1.261	[0.537, 1.996]	0.682	[0.589, 0.784]	0.658	[0.596, 0.719]
$100\sigma_z$	0.501	[0.218, 0.796]	0.180	[0.146, 0.213]	0.176	[0.144, 0.206]

Identification

- Likelihood contour from simulated data
- $\kappa = \frac{\tau}{\pi^2 \varphi} \left(\frac{1 - \nu}{\nu} \right)$



Expected Utility as Welfare Measures

- Infinite sum of the discount utilities

$$\begin{aligned} W_0 &= \mathbf{E}_0 \left[\sum_{t=0}^{\infty} \beta^t \left(\frac{c_t^{1-\tau} - 1}{1-\tau} - \chi_H y_t \right) \right] \\ &= \mathbf{E}_0 \left[\sum_{t=0}^{\infty} \beta^t \left(\frac{c_t^{1-\tau} - 1}{1-\tau} \right) \right] - \chi_H \mathbf{E}_0 \left[\sum_{t=0}^{\infty} \beta^t y_t \right] \\ &= U_0 - \chi_H V_0 \end{aligned}$$

- Approximations in recursive form

$$U_t = \frac{c_t^{1-\tau} - 1}{1-\tau} + \beta \mathbf{E}_t U_{t+1}, \quad V_t = y_t + \beta \mathbf{E}_t V_{t+1}$$

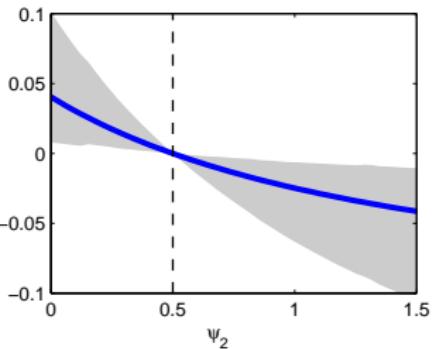
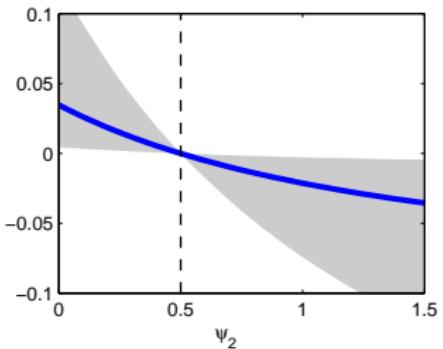
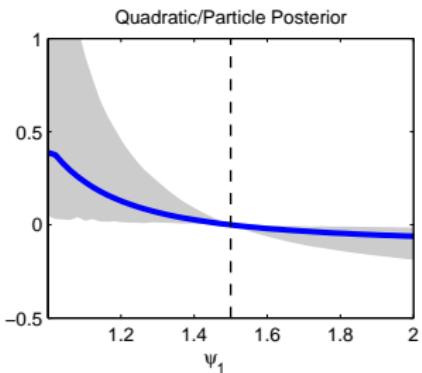
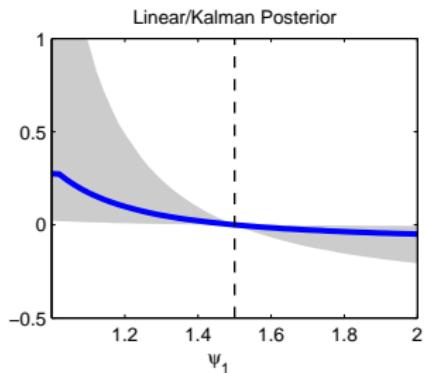
Welfare Cost

- Two sets of parameters, reference and alternative, implies different levels of welfare.
- The fraction of consumption in the alternative to be increased (or to be given up) to achieve the welfare level of the reference.

$$W_t^R = \mathbf{E}_t \left(\sum_{s=0}^{\infty} (\beta^A)^s \left(\frac{[(1 + \lambda) C_{t+s}^A / A_{t+s}^A]^{1-\tau^A} - 1}{1 - \tau^A} - \chi_H^A H_{t+s}^A \right) \right)$$

- Welfare differentials
 - Fix nonpolicy parameters
 - Differentials relative to baseline policy parameter

Welfare Cost



Generalized Impulse Responses

- Impulse responses in a linear model
 - Symmetry
 - Scalability
 - Path-independency: past and future
- Koop, Pesaran, and Potter (1996): Generalized Impulse Response

$$GI(n) = \mathbf{E}[x_{t+n}|\epsilon_t, \Omega_{t-1}] - \mathbf{E}[x_{t+n}|\Omega_{t-1}]$$

Generalized Impulse Responses

- Consider a nonlinear process

$$X_t = f(X_{t-1}) + \epsilon_t$$

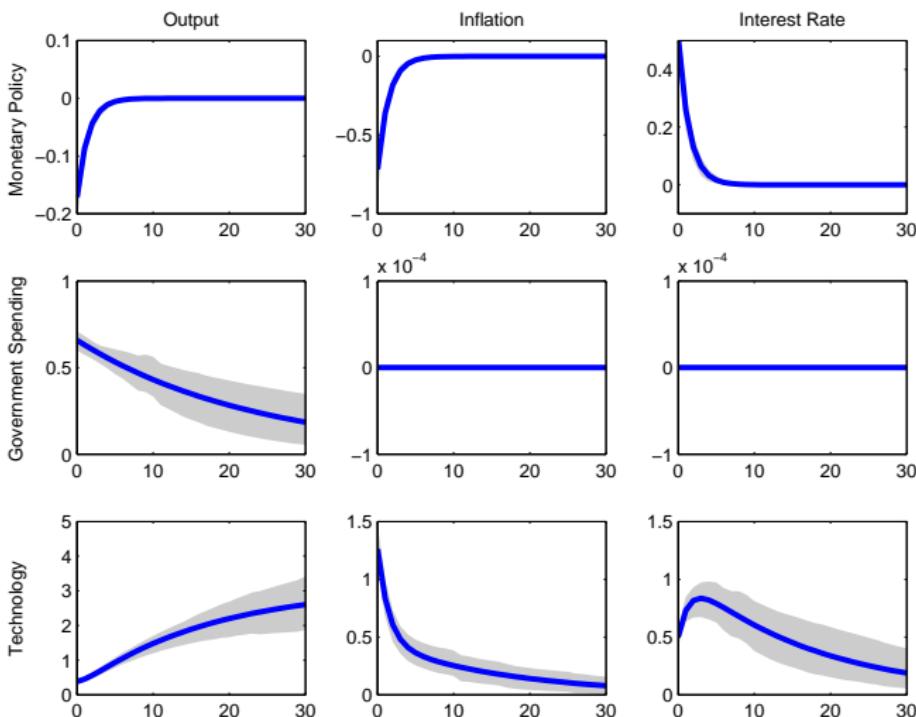
and two alternative shock processes which differ only in time t , that is, $\epsilon_k = \epsilon'_k$ for all $k \neq t$.

- Then conventional impulse response at time $t+2$ can be written as

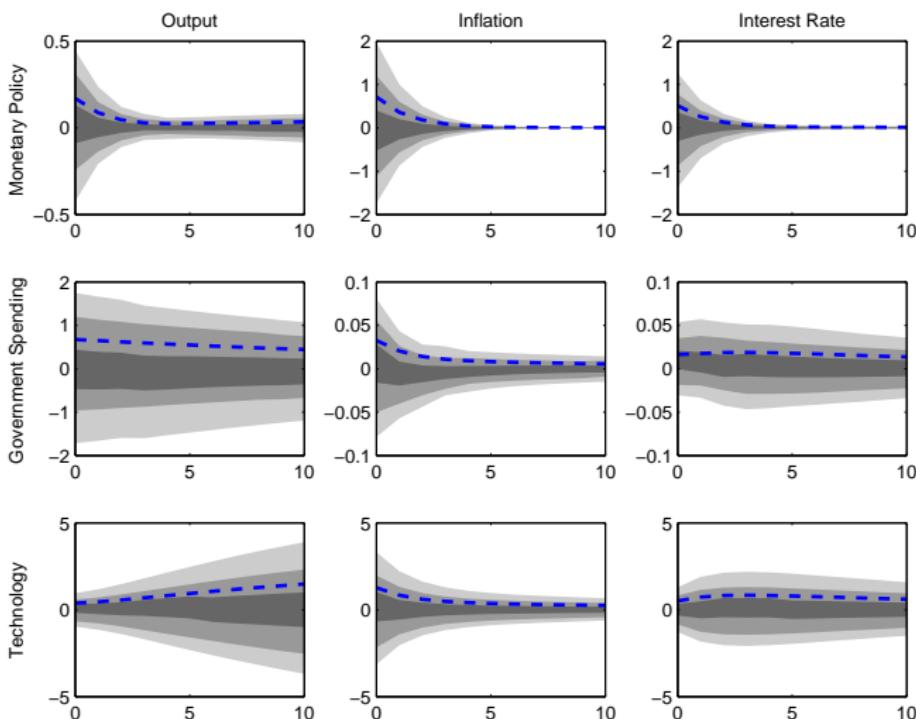
$$\begin{aligned} X'_{t+2} - X_{t+2} &= f[f\{f(X_{t-1}) + \epsilon_t\} + \epsilon_{t+1}] \\ &\quad - f[f\{f(X_{t-1}) + \epsilon'_t\} + \epsilon_{t+1}] \end{aligned}$$

which implies the path-dependence.

Impulse Responses



Impulse Responses For Nonlinear Processes



Conclusion

- Particle filter for second-order approximated DSGE model
- Stylized monetary model in line with Woodford (2003)
- U.S. quarterly data
- More identification on structural parameters
- Better fit as measured by the marginal data density
- Bias corrections on parameter estimates
- Aggressive estimates on welfare cost
- Unveils neglected impulse response dynamics